31/Mar/2012 International Workshop on Statistical Machine Learning for Speech Processing -Scalable Approach in the Era of Abundant Data-

Independent Component Analysis and Unsupervised Learning



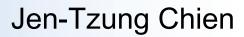


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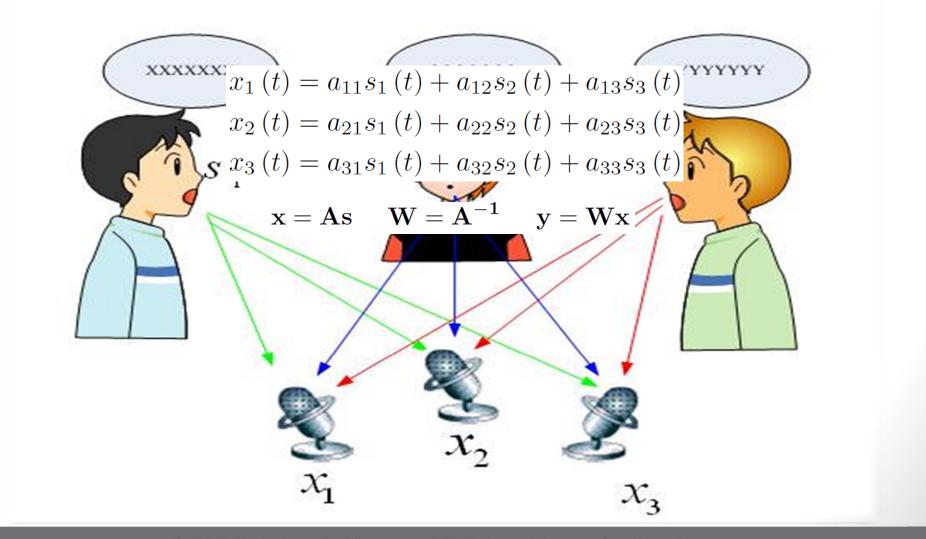
Signal and Information Processing

- 1. Independent Component Analysis
- 2. Case Study I: Speech Recognition
 - Independent voices
 - Nonparametric likelihood ratio ICA
- 3. Case Study II: Blind Source Separation
 - Convex divergence ICA
 - Nonstationary Bayesian ICA
 - Online Gaussian process ICA
- 4. Summary

Introduction

- Independent component analysis (ICA) is essential for blind source separation.
- ICA is applied to separate the *mixed signals* and find the *independent components*.
- The demixed components can be grouped into *clusters* where the *intra-cluster* elements are *dependent* and *inter-cluster* elements are *independent*.
- ICA provides *unsupervised learning* approach to acoustic modeling, signal separation and many others.

Blind Source Separation



APSIPA DL: Independent Component Analysis and Unsupervised Learning

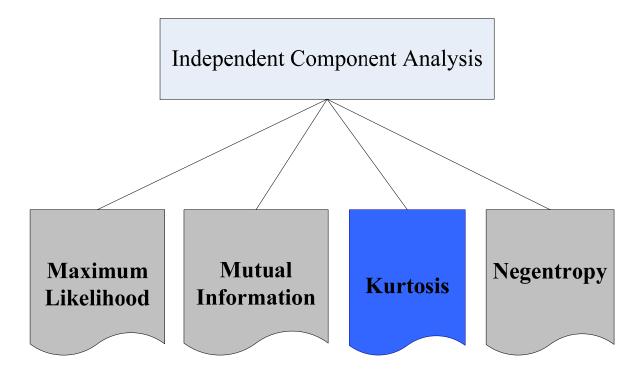
Independent Component Analysis

- Three assumptions
 - − sources ⇒ statistically independent
 - independent component is nongaussian distribution
 - − mixing matrix ⇒ square matrix

$$\mathbf{x} = \mathbf{A}\mathbf{s}$$

$$\begin{bmatrix} X_{11} & \dots & X_{1t} \\ \vdots & \ddots & \vdots \\ X_{m1} & \dots & X_{mt} \end{bmatrix} = \begin{bmatrix} A_{11} & \dots & A_{1m} \\ \vdots & \ddots & \vdots \\ A_{m1} & \dots & A_{mm} \end{bmatrix} \begin{bmatrix} S_{11} & \dots & S_{1t} \\ \vdots & \ddots & \vdots \\ S_{m1} & \dots & S_{mt} \end{bmatrix}$$
$$m \times t \qquad m \times m \qquad m \times t$$

ICA Objective Function



ICA Learning Rule

• ICA demixing matrix can be estimated by optimizing an *objective function* $D(\mathbf{X}, \mathbf{W})$ via *gradient descent* algorithm or *natural gradient* algorithm

$$\mathbf{W}^{(n+1)} = \mathbf{W}^{(n)} - \eta \frac{\partial D\left(\mathbf{X}, \mathbf{W}^{(n)}\right)}{\partial \mathbf{W}^{(n)}}$$

$$\mathbf{W}^{(n+1)} = \mathbf{W}^{(n)} - \eta \frac{\partial D\left(\mathbf{X}, \mathbf{W}^{(n)}\right)}{\partial \mathbf{W}^{(n)}} \mathbf{W}^{(n)\mathrm{T}} \mathbf{W}^{(n)}$$

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ICA for Speech Recognition

- Mismatch between training and test data always exists. *Adaptation* of HMM parameters is important.
- Eigenvoice (PCA) versus Independent Voice (ICA)

 PCA performs a linear de-correlation process
 ICA extracts the higher-order statistics

 $E[e_1e_2\cdots e_M] = E[e_1]E[e_2]\cdots E[e_M] \quad \text{uncorrelation } \Rightarrow \mathsf{PCA}$ $E[s_1^rs_2^r\cdots s_M^r] = E[s_1^r]E[s_2^r]\cdots E[s_M^r] \quad \begin{array}{l} \text{higher-order correlations} \\ \text{are zero } \Rightarrow \mathsf{ICA} \end{array}$

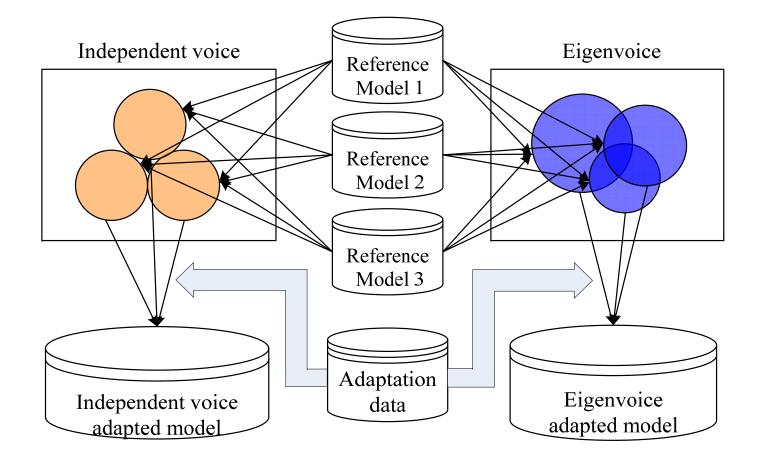
Sparseness & Information Redundancy

- The degree of *sparseness* in distribution of the transformed signals is proportional to the amount of *information* conveyed by the transformation.
- Sparseness measurement

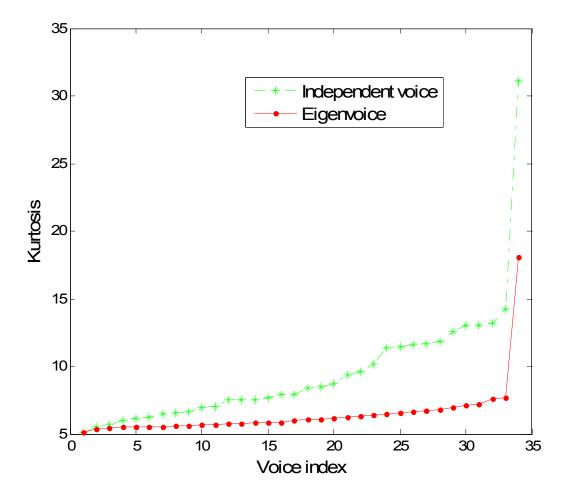
- fourth-order statistics (*kurtosis*) \Rightarrow nongaussianity kurt(s) = $E[s^4]/E^2[s^2] - 3$

 Information redundancy reduction using ICA is higher than that using PCA.

Eigenvoices versus Independent Voices



Evaluation of Kurtosis



Word Error Rates on Aurora2

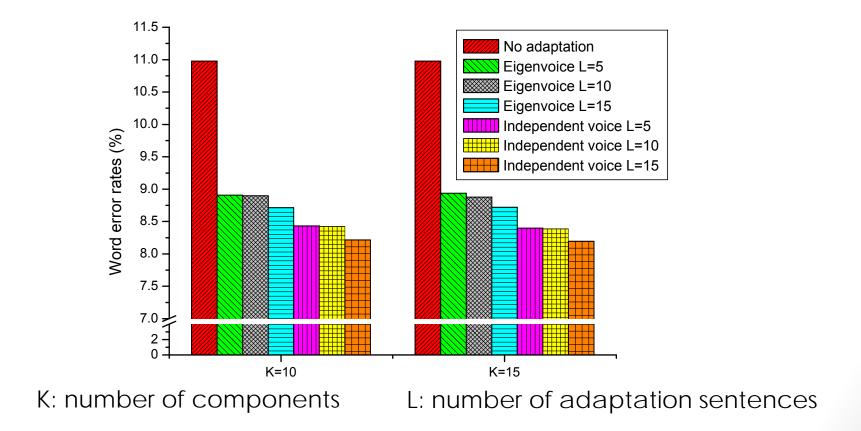


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Test of Independence

- Given the demixing signals y = Wx, the null & alternative hypotheses are defined as $H_0: y^{(1)}, y^{(2)}, \ldots, y^{(d)}$ are mutually independent $H_1: y^{(1)}, y^{(2)}, \ldots, y^{(d)}$ are not mutually independent.
- If y is *Gaussian* distributed, we are testing whether the correlation between $y^{(i)}$ and $y^{(j)}$ is equal to zero, i.e. $\Sigma_D = \text{diag}\{\sigma_{ii}^2\}$ or

$$H_0: \sigma_{ij}^2 = E\left\{ \left(y^{(i)} - \mu^{(i)} \right) \left(y^{(j)} - \mu^{(j)} \right) \right\} = 0, \quad \text{for all } i \neq j.$$

Likelihood Ratio

$$\lambda_{\rm LR} = \frac{p(Y|H_0)}{p(Y|H_1)} = \frac{\max_{\boldsymbol{\mu}, \Sigma_D} p(Y|\boldsymbol{\mu}, \Sigma_D)}{\max_{\boldsymbol{\mu}, \Sigma} p(Y|\boldsymbol{\mu}, \Sigma)}$$

- LR serves as the *test statistics* which measures the confidence for H_0 against H_1 .
- LR is a measure of *independence* for
 Y = {y₁, y₂,..., y_T} and can act as an objective function for finding ICA demixing matrix.
- However, it is *not allowed* to assume Gaussianity for ICA problem.

Nonparametric Approach

- Let each sample be transformed by $\mathbf{y}_t = W \mathbf{x}_t$.
- Instead of assuming Gaussianity, we apply the kernel density estimation

$$p\left(y^{(i)}\right) = \frac{1}{Th} \sum_{t=1}^{T} \varphi\left(\frac{y^{(i)} - y_t^{(i)}}{h}\right), \quad i = 1, \dots, d$$

using Gaussian kernel

$$\varphi(u) = \frac{1}{\sqrt{2\pi}} e^{-\frac{u^2}{2}}$$

• Kernel centroid is given by $y_t^{(i)} = \mathbf{w}_i \mathbf{x}_t = \sum_{j=1}^d w_{ij} x_t^{(j)}$

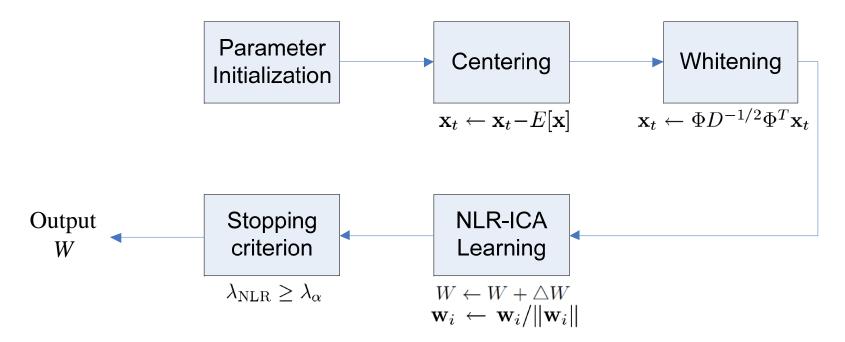
Nonparametric Likelihood Ratio

• NLR objective function

$$\lambda_{\text{NLR}} = \frac{p(Y|H_0)}{p(Y|H_1)} = \frac{\prod_{t=1}^T \prod_{i=1}^d p\left(y_k^{(i)}\right)}{\prod_{t=1}^T p\left(\mathbf{y}_k\right)}$$
$$= \frac{\prod_{t=1}^T \prod_{i=1}^d \left[\frac{1}{Th} \sum_{k=1}^T \varphi\left(\frac{y_t^{(i)} - y_k^{(i)}}{h}\right)\right]}{\prod_{t=1}^T \left[\frac{1}{Th^d} \sum_{k=1}^T \psi\left(\frac{\mathbf{y}_t - \mathbf{y}_k}{h}\right)\right]}$$

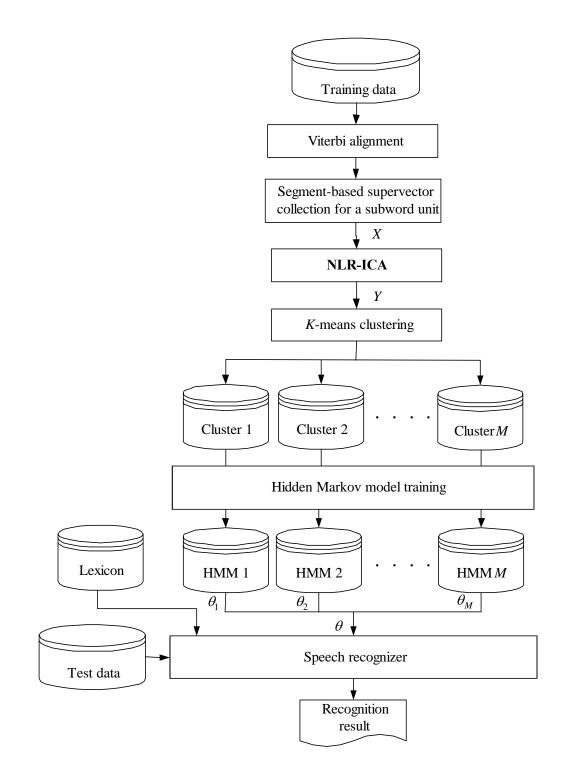
with multivariate Gaussian kernel $\psi(\mathbf{v}) = \frac{1}{\sqrt{(2\pi)^d}} e^{-\frac{1}{2}\mathbf{v}^T\mathbf{v}}$

ICA Learning Procedure



- Log likelihood ratio for null and alternative hypotheses $\log \lambda_{\rm NLR} = L_0(W) L_1(W)$
- Maximizing $\log \lambda_{\rm NLR}$ with respect to $W_{\rm NLR}$, we obtain

$$\Delta W = -\eta \left(\nabla_W L_0(W) - \nabla_W L_1(W) \right)$$



Segment-Based Supervector

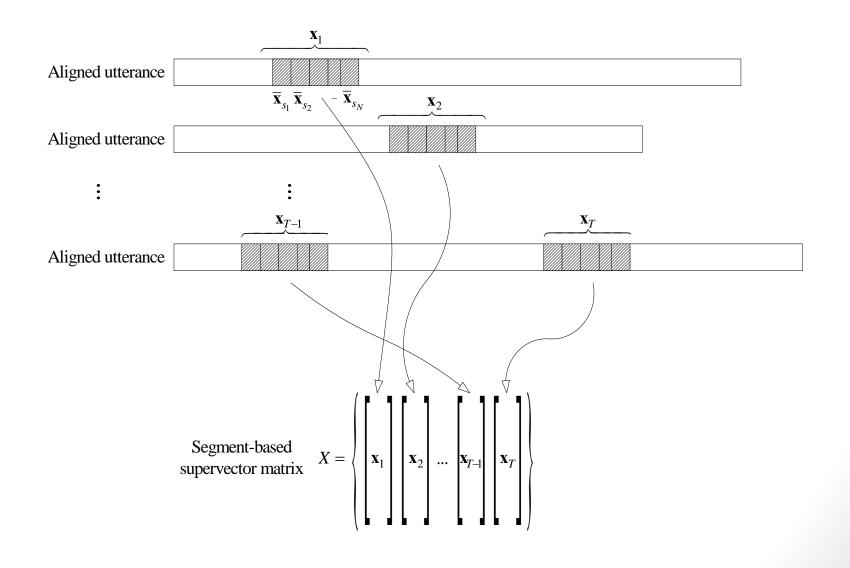


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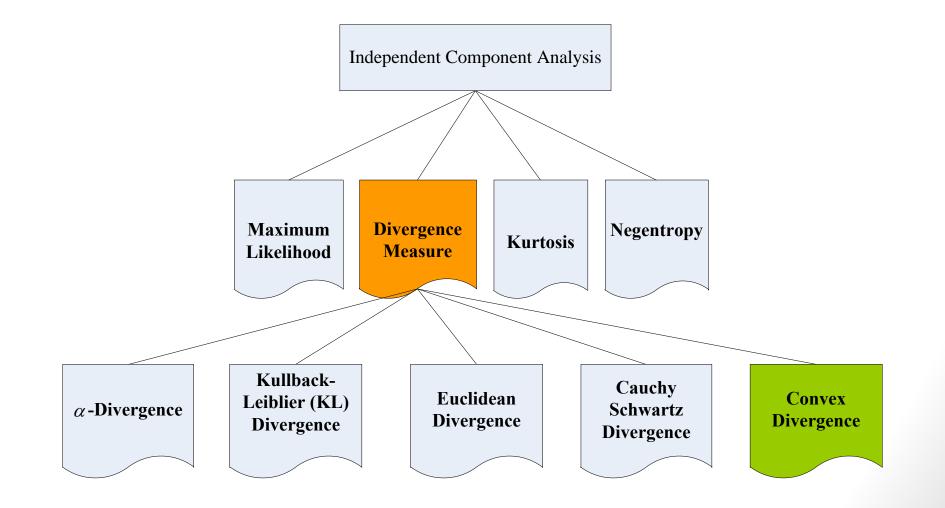
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ICA Objective Function



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Mutual Information & KL Divergence

- Mutual information between two variables y_1 and y_2 is defined by using the Shannon entropy $H[\cdot]$.
- It can be formulated as the KL divergence or *relative entropy* between the joint distribution $p(y_1, y_2)$ and the product of marginal distribution $p(y_1)p(y_2)$

$$D_{\mathrm{KL}}(y_1, y_2) = H[p(y_1)] + H[p(y_2)] - H[p(y_1, y_2)]$$
$$= \iint p(y_1, y_2) \log \frac{p(y_1, y_2)}{p(y_1)p(y_2)} dy_1 dy_2$$

where $D_{\mathrm{KL}}(y_1, y_2) \geq 0$

Divergence Measures

• Euclidean divergence $D_{\rm E}(y_1, y_2)$

 $\iint (p(y_1, y_2) - p(y_1)p(y_2))^2 dy_1 dy_2$

• Cauchy-Schwartz divergence $D_{CS}(y_1, y_2)$

$$\log\left\{\frac{\iint p(y_1, y_2)^2 dy_1 dy_2 \cdot \iint p(y_1)^2 p(y_2)^2 dy_1 dy_2}{\left[\iint p(y_1, y_2) p(y_1) p(y_2) dy_1 dy_2\right]^2}\right\}$$

• α -divergence $D_{\alpha}(y_1, y_2, \alpha)$

$$\frac{4}{1-\alpha^2} \iint \left[\frac{1-\alpha}{2} p(y_1, y_2) + \frac{1+\alpha}{2} p(y_1) p(y_2) - p(y_1, y_2)^{(1-\alpha)/2} (p(y_1) p(y_2))^{(1+\alpha)/2} \right] dy_1 dy_2$$

Divergence Measures

• f-divergence $D_f(y_1, y_2)$

$$\iint p(y_1)p(y_2)f\left(\frac{p(y_1,y_2)}{p(y_1)p(y_2)}\right)dy_1dy_2$$

• Jensen-Shannon divergence $D_{JS}(y_1, y_2)$ $H[\lambda p(y_1, y_2) + (1 - \lambda)p(y_1)p(y_2)]$ $-\lambda H[p(y_1, y_2)] - (1 - \lambda)H[p(y_1)p(y_2)]$

where $0 \le \lambda \le 1$. Entropy is a concave function.

Convex Function

• A convex function should meet the *Jensen's inequality*

$$f(\lambda p(y_1, y_2) + (1 - \lambda)p(y_1)p(y_2)) \\\leq \lambda f(p(y_1, y_2)) + (1 - \lambda)f(p(y_1)p(y_2))$$

• A general convex function is defined by

$$f(t) = \frac{4}{1 - \alpha^2} \left[\frac{1 - \alpha}{2} + \frac{1 + \alpha}{2} t - t^{(1 + \alpha)/2} \right]$$

$$f(t) \ge 0$$
 for $t \ge 0, f(1) = 0$ and $f'(1) = 0$.

Convex Divergence

• By assuming equal weight $\lambda = 1/2$, we have

$$\frac{2}{1-\alpha^2} \iint \left\{ 2 \left[\frac{p(y_1, y_2) + p(y_1)p(y_2)}{2} \right]^{(1+\alpha)/2} - \left[p(y_1, y_2)^{(1+\alpha)/2} + (p(y_1)p(y_2))^{(1+\alpha)/2} \right] \right\} dy_1 dy_2.$$

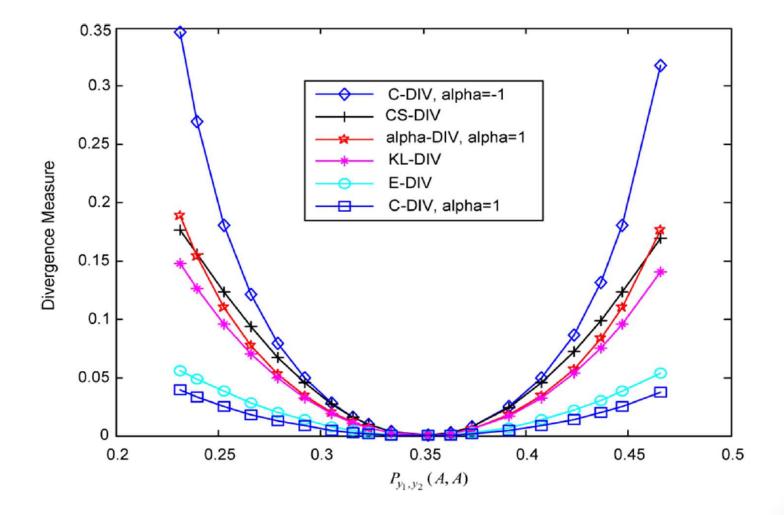
• When $\alpha = -1$, C-DIV is derived as a case with convex function $-\log(\cdot)$

$$D_{\mathcal{C}}(y_1, y_2, -1) = \iint \left\{ \log \left[\frac{p(y_1, y_2) + p(y_1)p(y_2)}{2} \right] -\frac{1}{2} \log(p(y_1, y_2)) - \frac{1}{2} \log(p(y_1)p(y_2)) \right\} dy_1 dy_2$$

Different Divergence Measures

	Symmetric	Convexity	Combination	Special Realization from
	Divergence	Parameter	Weight	other DIV
KL-DIV	No	no	No	α -DIV ($\alpha = -1$)
α-DIV	No	α	No	<i>f</i> -DIV (using (6))
<i>f</i> -DIV	No	α	No	no
JS-DIV	$\lambda = 0.5$	no	λ (convex)	C-DIV ($\lambda = 0.5, \alpha = 1$)
Convex-log DIV	$\lambda = 0.5$	no	λ (convex)	C-DIV ($\lambda = 0.5, \alpha = -1$)
C-DIV	$\lambda = 0.5$	α	λ (convex)	no

Different Divergence Measures



Convex Divergence ICA

C-ICA learning algorithm

$$\mathbf{W}^{(n+1)} = \mathbf{W}^{(n)} - \eta \frac{\partial D\left(\mathbf{X}, \mathbf{W}^{(n)}\right)}{\partial \mathbf{W}^{(n)}}$$

$$D_{\mathrm{C}}(\mathbf{X}, \mathbf{W}, \alpha) = \frac{2}{1 - \alpha^2} \sum_{t=1}^{T} \left\{ 2 \left[\frac{1}{2} \left(p(\mathbf{W}\mathbf{x}_t) + \prod_{m=1}^{M} p(\mathbf{w}_m \mathbf{x}_t) \right) \right]^{(1+\alpha)/2} - \left[p(\mathbf{W}\mathbf{x}_t)^{(1+\alpha)/2} + \left(\prod_{m=1}^{M} p(\mathbf{w}_m \mathbf{x}_t) \right)^{(1+\alpha)/2} \right] \right\}$$

• *Nonparametric* C-ICA is established by using Parzen window density function.

Simulated Experiments

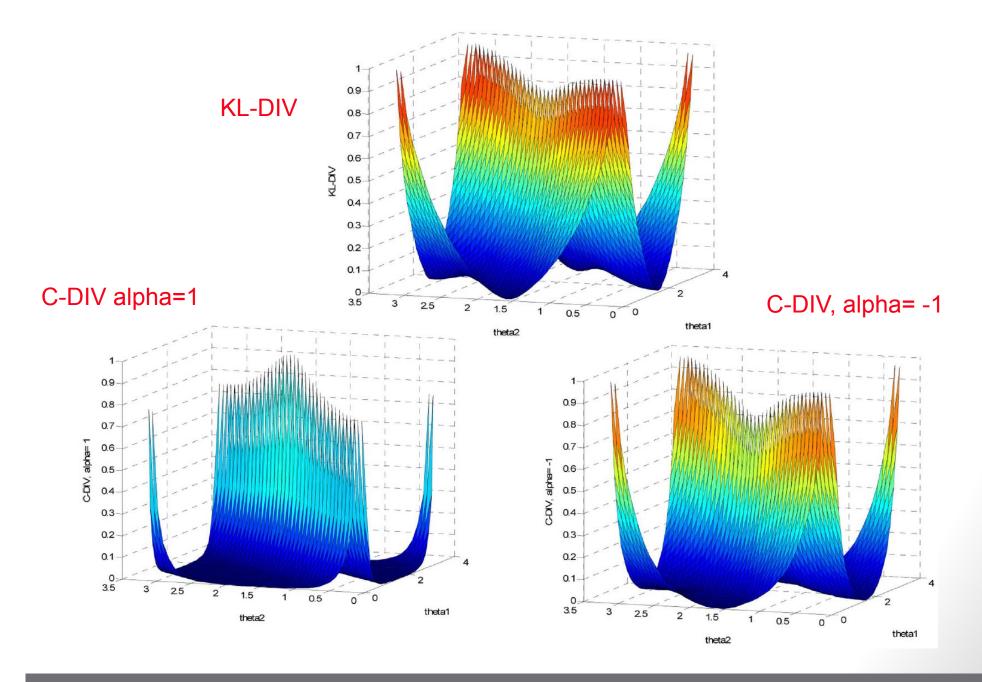
• A parametric demixing matrix

$$W = \begin{bmatrix} \cos \theta_1 & \sin \theta_1 \\ \cos \theta_2 & \sin \theta_2 \end{bmatrix}$$

• Two sources: *super-Gaussian* and *sub-Gaussian* distribution

$$p(s_1) = \begin{cases} \frac{1}{2\tau_1}, \ s_1 \in [-\tau_1, \tau_1] \\ 0, \ \text{otherwise} \end{cases} \quad p(s_2) = \frac{1}{2\tau_2} \exp\left(-\frac{|s_2|}{\tau_2}\right)$$

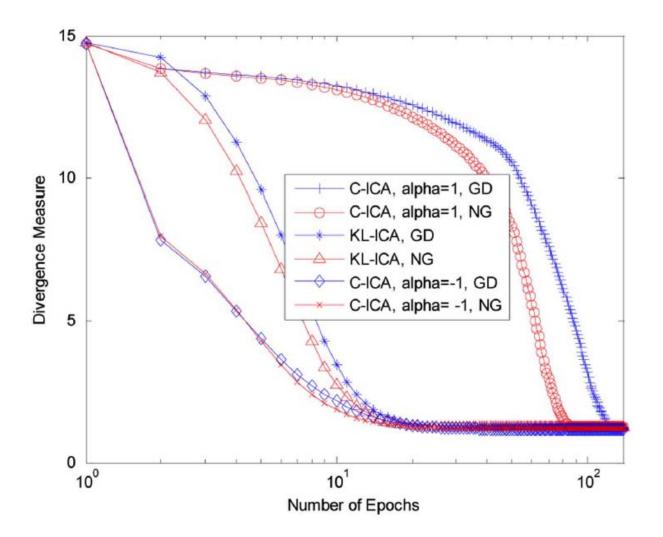
- Kurtosis
 - -Source 1: -1.13, source 2: 2.23



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Learning Curves



Experiments on Blind Source Separation

- One music signal and two speech signals from two male speakers were sampled from ICA'99 BSS Test Sets at http://sound.media.mit.edu/ica-bench/
- Mixing matrix

$$A = \begin{bmatrix} 0.8 & 0.2 & 0.3 \\ 0.3 & -0.8 & 0.2 \\ -0.3 & 0.7 & 0.3 \end{bmatrix}$$

Evaluation metric

-signal-to-interference ratio (SIR) SIR(dB) = $10\log_{10}\left(\sum_{t=1}^{T} \|\mathbf{s}_{t}\|^{2} / \sum_{t=1}^{T} \|\mathbf{y}_{t} - \mathbf{s}_{t}\|^{2}\right)$

Comparison of Different Methods

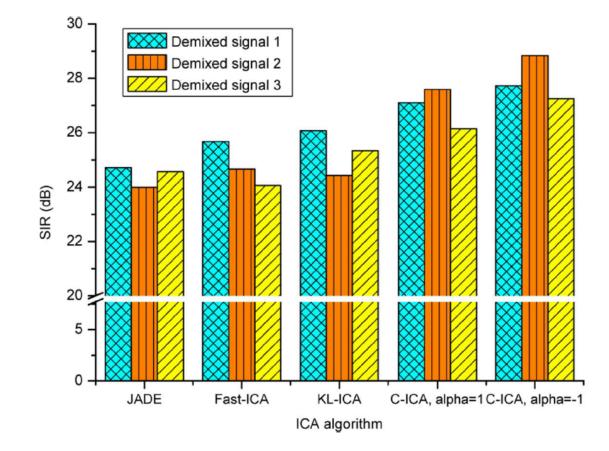


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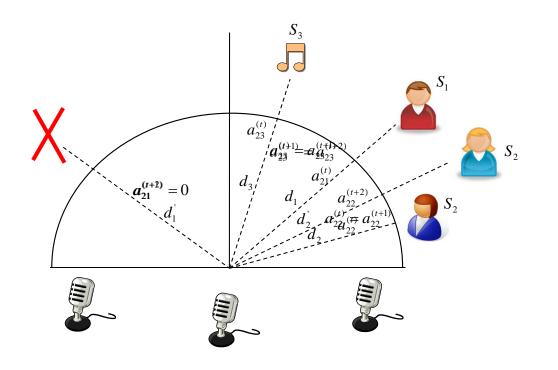


Why Nonstationary Source Separation?

- Real-world *blind source separation*
 - -number of sources is unknown
 - -BSS is a dynamic time-varying system
 - mixing process is nonstationary
- Why nonstationary?
 - Bayesian method using ARD can determine the changing number of sources
 - recursive Bayesian for online tracking of nonstationary conditions
 - Gaussian process provides a nonparametric solution to represent temporal structure of time-varying mixing system.

Nonstationary Mixing Systems

- Time-varying mixing matrix
- Source signals may abruptly appear or disappear



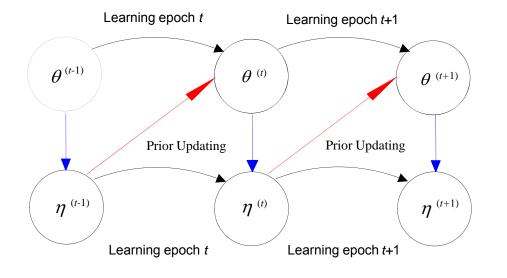
Nonstationary Bayesian (NB) Learning

 Maximum a posteriori estimation of NB-ICA parameters and compensation parameters

$$\theta^{(t)} = \arg \max_{\theta} p(X_t | \theta^{(t)}, \eta^{(t-1)}) p(\theta)$$

$$\eta^{(t)} = \arg \max_{\eta} p(X_t | \theta^{(t)}, \eta) p(\eta | \phi^{t-1})$$

updating



Model Construction

- Noisy ICA model $\mathbf{x}_t = A\mathbf{s}_t + \boldsymbol{\varepsilon}_t$
- Likelihood function of an observation \mathbf{x}_t $p(\mathbf{x}_t | \mathbf{A}^{(t)}, \mathbf{s}^{(t)}, \beta^{(t)}) = N(\mathbf{x}_t | \mathbf{A}^{(t)} \mathbf{s}^{(t)}, \beta^{(t)^{-1}} I_N)$
- Distribution of model parameters - source $p(\mathbf{s}^{(t)}|\boldsymbol{\pi}^{(t)}, \boldsymbol{\mu}^{(t)}, \boldsymbol{\gamma}^{(t)}) = \prod_{m=1}^{M} \left[\sum_{k=1}^{K} \pi_k^{(t)} N(s_m^{(t)}|\boldsymbol{\mu}_k^{(t)}, \boldsymbol{\gamma}_k^{(t)^{-1}}) \right]$ - mixing matrix $p(A^{(t)}|\boldsymbol{\alpha}^{(t)}) = \prod_{m=1}^{M} \left[\prod_{n=1}^{N} N(a_{nm}^{(t)}|0, \boldsymbol{\alpha}_m^{(t)^{-1}}) \right]$

-noise
$$p(\varepsilon_t|\beta^{(t)}) = N(\varepsilon_t|0, \beta^{(t)^{-1}}I_N)$$

Prior & Marginal Distributions

Prior distributions

-precision of noise $p(\beta^{(t)}|u_{\beta}, w_{\beta}) = \text{Gam}(\beta^{(t)}|u_{\beta}, w_{\beta})$ -precision of mixing matrix $p(\boldsymbol{\alpha}^{(t)}|u_{\alpha}, w_{\alpha}) = \prod_{m=1}^{M} \text{Gam}(\alpha_{m}^{(t)}|u_{\alpha}, w_{\alpha})$

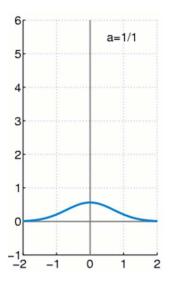
• Marginal likelihood of NB-ICA model

$$p(X) = \prod_{t=1}^{T} \int p(\mathbf{x}_t | A^{(t)}, \mathbf{s}^{(t)}, \boldsymbol{\alpha}^{(t)}, \beta^{(t)}) p(A^{(t)} | \boldsymbol{\alpha}^{(t)}) p(\boldsymbol{\alpha}^{(t)} | u_{\alpha}^{(t)}, w_{\alpha}^{(t)})$$
$$\times p(\mathbf{s}^{(t)} | \boldsymbol{\pi}^{(t)}, \boldsymbol{\mu}^{(t)}, \boldsymbol{\gamma}^{(t)}) p(\boldsymbol{\beta}^{(t)} | \boldsymbol{u}_{\beta}^{(t)}, \mathbf{w}_{\beta}^{(t)}) dA^{(t)} d\mathbf{s}^{(t)} d\boldsymbol{\alpha}^{(t)} d\boldsymbol{\beta}^{(t)}$$

Automatic Relevance Determination

• Detection of source signals

$$\alpha_m^{(t)} = \begin{cases} \infty & , \ a_m^{(t)} = \{a_{nm}^{(t)}\} \to 0\\ <\infty & , \ a_m^{(t)} = \{a_{nm}^{(t)}\} \neq 0 \end{cases}$$



- number of sources can be determined

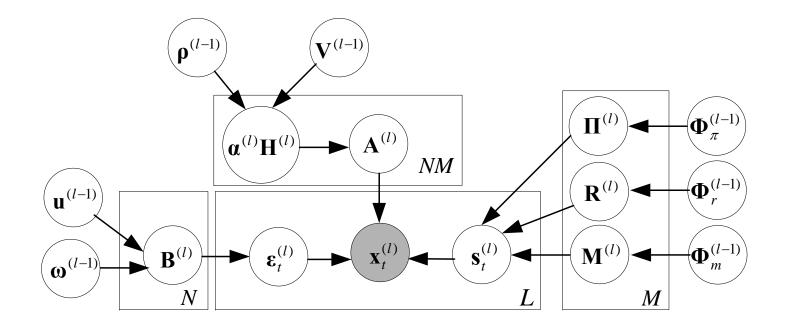
Compensation for Nonstationary ICA

 $G_{\eta^{(t)}}(\alpha^{(t)}) = \eta^{(t)}\alpha^{(t)}$

- Prior density of compensation parameter
 - conjugate prior (Wishart distribution)

$$p(\eta^{(t)}|\varphi^{t-1} = \{\nu^{t-1}, \Lambda^{t-1}\})c(M, \nu^{t-1}) \left|\frac{1}{2}\nu^{t-1}\Lambda^{t-1}\right|^{(\nu^{t-1}-1)/2} \\ \times \left|\eta^{(t)}\right|^{(\nu^{t-1}-M-2)/2} \exp\left[-\frac{1}{2}\operatorname{tr}\left(\nu^{t-1}\Lambda^{t-1}\eta^{(t)}\right)\right] \\ c(M, \nu^{t-1}) = \left(\pi^{M(M-1)/4}\prod_{m=1}^{M}\Gamma\left(\left(\nu^{t-1}-m\right)/2\right)\right)^{-1}$$

Graphical Model for NB-ICA

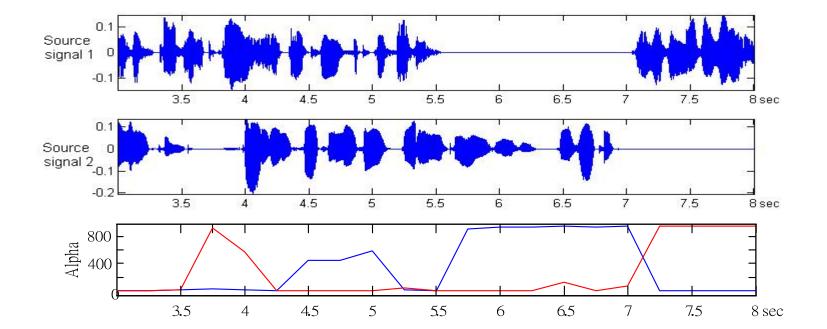


Experiments

- Nonstationary Blind Source Separation
 ICA'99 http://sound.media.mit.edu/ica-bench/
- Scenarios
 - -state of source signals: active or inactive
 - source signals or sensors are moving: nonstationary mixing matrix

$$A^{(t)} = \begin{bmatrix} \cos(2\pi f_1 t) & \sin(2\pi f_2 t) \\ -\sin(2\pi f_1 t) & \cos(2\pi f_2 t) \end{bmatrix}$$
$$f_1 = 1/5 \text{ Hz} \quad f_2 = 1/2.5 \text{ Hz}$$

Source Signals and ARD Curves



Blue: first source signal Red: second source signal

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Online Gaussian Process (OLGP)

- Basic ideas
 - incrementally detect the status of source signals and estimate the corresponding distributions from online observation data $\chi^t = \left\{ X^{(1)}, X^{(2)}, \cdots, X^{(t)} \right\}.$
 - *temporal structure* of time-varying mixing coefficients $A^{(t)}$ are characterized by *Gaussian process*.
 - Gaussian process is a nonparametric model which defines the prior distribution over functions for Bayesian inference.

Model Construction

- Noisy ICA model $\mathbf{x}^{(t)} = A^{(t)}\mathbf{s}^{(t)} + \boldsymbol{\varepsilon}^{(t)}$
- Likelihood function

$$p(\mathbf{x}^{(t,i)} \mid \mathbf{A}^{(t,i)}, \mathbf{s}^{(t,i)}, \beta^{(t,i)}) = N(\mathbf{x}^{(t,i)} \mid \mathbf{A}^{(t,i)}\mathbf{s}^{(t,i)}, \beta^{(t,i)^{-1}}I_N)$$

• Distribution of model parameters

- source

$$p(\mathbf{s}^{(t,i)} \mid \boldsymbol{\pi}^{(t,i)}, \boldsymbol{\mu}^{(t,i)}, \boldsymbol{\gamma}^{(t,i)}) = \prod_{m=1}^{M} \left[\sum_{k=1}^{K} \pi_{m,k}^{(t,i)} N(s_{m}^{(t,i)} \mid \mu_{m,k}^{(t,i)}, \gamma_{m,k}^{(t,i)^{-1}}) \right]$$

- noise

$$p(\varepsilon^{(t,i)} \mid \beta) = N(\varepsilon^{(t,i)} \mid 0, \beta_m^{(t,i)^{-1}} I_N)$$

-
$$p(\beta^{(t,i)}) = \text{Gam}(u_{\beta}^{(t,i)}, w_{\beta}^{(t,i)})$$

Gaussian Process

• Mixing matrix

- $A^{(t)}$ is generated by the latent function $f_{nm}(\cdot)$

$$a_{nm}^{(t,i)} = f_{nm} \left(\mathbf{a}_{nm}^{(t,i-1,i-p)} \right)$$
$$\mathbf{a}_{nm}^{(t,i-1,i-p)} = \left[a_{nm}^{(t,i-1p)} a_{nm}^{(t,i-2)} \cdots a_{nm}^{(t,i-p)} \right]^{T}$$

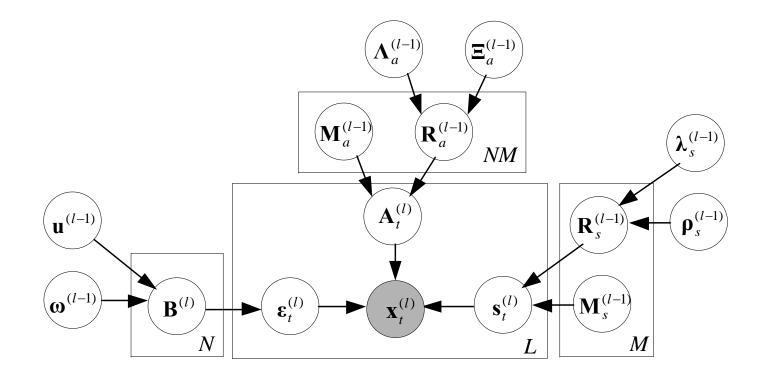
- GP is adopted to describe the distribution of $f_{nm}(\cdot)$

$$f_{nm}(\mathbf{a}_{nm}^{(t,i-1,i-p)}) \sim N(0,\kappa(\mathbf{a}_{nm}^{(t,i-1,i-p)},\mathbf{a}_{nm}^{(t,\tau-1,\tau-p)}))$$

$$\kappa(\mathbf{a}_{nm}^{(t,i-1,i-p)},\mathbf{a}_{nm}^{(t,\tau-1,\tau-p)}) = \rho_{nm}^{(t,i)} \exp\left\{-\frac{\lambda_{nm}^{(t,i)}}{2} \left\|\mathbf{a}_{nm}^{(t,i-1,i-p)} - \mathbf{a}_{nm}^{(t,\tau-1,\tau-p)}\right\|^{2}\right\}$$

- $\left\{\lambda_{nm}^{(t,i)}, \rho_{nm}^{(t,i)}\right\}$ are *hyperparameters* of kernel function

Graphical Model for OLGP-ICA

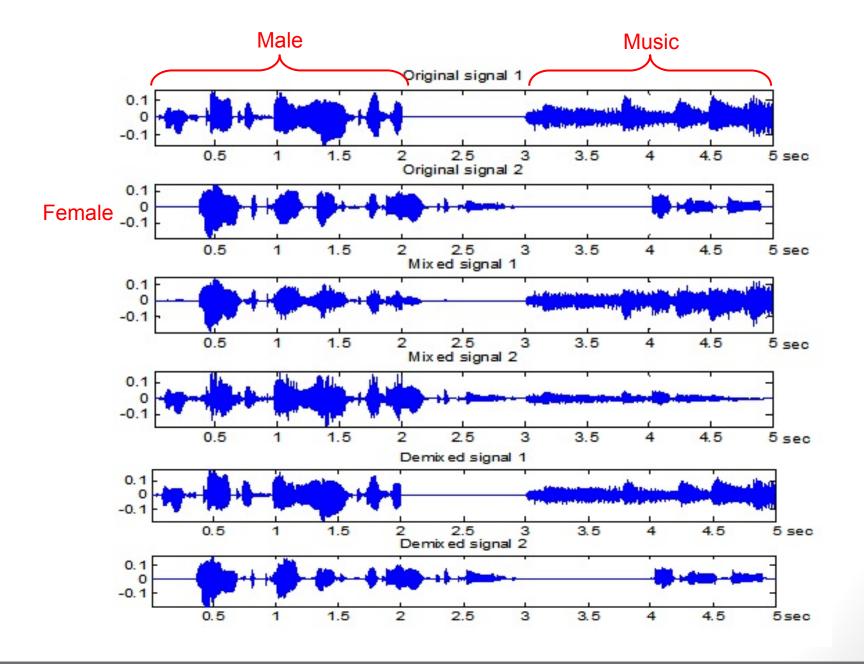


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Experimental Setup

- Nonstationary source separation using source signals from – http://www.kecl.ntt.co.jp/icl/signal/
- Nonstationary scenarios
 - -status of source signals: active or inactive
 - source signals or sensors are moving: nonstationary mixing matrix

$$A^{(t)} = \begin{bmatrix} \cos(2\pi f_1 t) & \sin(2\pi f_2 t) \\ -\sin(2\pi f_1 t) & \cos(2\pi f_2 t) \end{bmatrix}$$
$$f_1 = 1/20 \text{ Hz} \qquad f_2 = 1/10 \text{ Hz}$$



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Comparison of Different Methods

• Signal-to-interference ratios (SIRs) (dB)

	VB-ICA	BICA-HMM	Switching- ICA	Online VB-ICA	OLGP-ICA
Demixed signal 1	7.97	9.04	12.06	11.26	17.24
Demixed signal 2	-3.23	-1.5	-4.82	4.47	9.96

Summary

- We presented speaker adaptation method based on independent voices by fulfilling ICA perspective.
- A *nonparametric likelihood ratio* ICA was proposed according to hypothesis test theory.
- A *convex divergence* was developed as an optimization metric for ICA algorithm.
- A *nonstationary Bayesian* ICA was proposed to deal with nonstationary mixing system.
- An online Gaussian process ICA was presented for nonstationary and temporally correlated source separation.
- ICA methods could be extended to solve *nonnegative matrix factorization* and single-channel separation.

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Thanks to



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